

Strong Consistency of M Estimator in Linear Model for $\tilde{\varphi}$ -mixing Samples*

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Abstract: The strong consistency of M estimator of regression parameter in linear model for $\tilde{\varphi}$ -mixing samples is discussed by using the classic Rosenthal type inequality. We get the strong consistency of M estimator under lower moment condition, which generalizes and improves the corresponding ones for independent sequences.

Key words: $\tilde{\varphi}$ -mixing sample, M estimator, strong consistency

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1 Introduction

Consider the following linear model:

$$y_i = \mathbf{x}'_i \boldsymbol{\beta}_0 + e_i, \quad i = 1, 2, \dots, n, \quad n \in \mathbf{N}, \quad (1.1)$$

where \mathbf{x}_i ($i = 1, 2, \dots, n$) is a known p -dimensional vector, $\boldsymbol{\beta}_0$ is an unknown p -dimensional regression parametric vector, and e_1, e_2, \dots, e_n are random errors. Let f be a convex function on \mathbf{R} . The M estimator of $\boldsymbol{\beta}_0$ is $\hat{\boldsymbol{\beta}}_n$ satisfying the following equation:

$$\sum_{i=1}^n f(y_i - \mathbf{x}'_i \hat{\boldsymbol{\beta}}_n) = \min_{\boldsymbol{\beta} \in \mathbf{R}^p} \sum_{i=1}^n f(y_i - \mathbf{x}'_i \boldsymbol{\beta}). \quad (1.2)$$

The scope of M estimator is very wide containing the least squares estimator, maximum likelihood estimator *etc.* Since Huber^[1] studied the M estimator of regression parameter in linear model, many authors have shown great interest in this field and obtained many useful results; see, for example, [2–5].

Throughout this paper, we use the following notations: Let

$$\boldsymbol{\alpha} = (a_1, a_2, \dots, a_p)'$$

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be a p -dimensional vector, and

$$\|\boldsymbol{\alpha}\|^2 \doteq \sum_{i=1}^p a_i^2 = \boldsymbol{\alpha}'\boldsymbol{\alpha}, \quad |\boldsymbol{\alpha}| \doteq \max_{1 \leq i \leq p} |a_i|.$$

Denote

$$\mathbf{S}_n = \sum_{i=1}^n \mathbf{x}_i \mathbf{x}_i'.$$

Assume that \mathbf{S}_n^{-1} exists, and

$$d_n \doteq \max_{1 \leq i \leq n} \mathbf{x}_i' \mathbf{S}_n^{-1} \mathbf{x}_i.$$

Let f be a convex function, ψ_- and ψ_+ denote the left derivative and right derivative of function f , respectively. $a_n = O(b_n)$ denotes that there exists a positive constant C such that $\left| \frac{a_n}{b_n} \right| \leq C$ for all sufficiently large n . C and C_i ($i \geq 1$) are positive constants which may be different in various places.

As for the sufficient condition for the strong consistency of M estimator, Chen and Zhao^[2] obtained the following result:

Theorem 1.1^[2] *Let $e_1, e_2, \dots, e_n, \dots$ be a sequence of independent random variables with identical distribution, and f be a convex function satisfying the following two conditions:*

- (1) *There exist constants $l_1 > 0$ and $l_2 > 0$ such that*

$$E(f(e_1 + u) - f(e_1)) \geq l_1 u^2, \quad |u| < l_2;$$

- (2) *There exists a constant $\Delta > 0$ such that*

$$E|\psi_+(e_1 \pm \Delta)|^m \leq h_m < \infty, \quad m = 1, 2, \dots$$

If there exists a δ with $0 < \delta \leq 1$ such that $d_n = O(n^{-\delta})$, then, $\hat{\boldsymbol{\beta}}_n$ is the strong consistency estimator of $\boldsymbol{\beta}_0$.

Yang^[3] improved the result of Theorem 1.1 and obtained the following Theorem 1.2 and Theorem 1.3.

Theorem 1.2^[3] *Let $e_1, e_2, \dots, e_n, \dots$ be a sequence of independent random variables with identical distribution, and f be a convex function satisfying the following two conditions:*

- (1) *There exist constants $l_1 > 0$ and $l_2 > 0$ such that*

$$E(f(e_1 + u) - f(e_1)) \geq l_1 u^2, \quad |u| < l_2;$$

- (2) *There exist constants $h_0 > 0$, $\Delta > 0$ and $0 < \delta \leq 1$ such that*

$$E|\psi_+(e_1 \pm \Delta)|^{2/\delta} \leq h_0 < \infty$$

and

$$d_n = O(n^{-\delta}).$$

Then, $\hat{\boldsymbol{\beta}}_n$ is the strong consistency estimator of $\boldsymbol{\beta}_0$.

Theorem 1.3^[3] *Let $e_1, e_2, \dots, e_n, \dots$ be a sequence of independent random variables, and f be a convex function satisfying the following two conditions:*